Bocconi



Contraction rates for conjugate gradient and Lanczos approximate posteriors in Gaussian process regression

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Joint work with



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Gaussian process (GP) regression

Consider i.i.d. observations from the model

$$Y_i = F(X_i) + \varepsilon_i, \qquad i = 1, \dots, n, \tag{1}$$

where

• $X_1, \ldots, X_n \sim G$ i.i.d. on \mathbb{R}^d and $\varepsilon \sim N(0, \sigma^2 I_n)$;

▶ $F \sim GP(0, k)$ with p.s.d. kernel $k : \mathbb{R}^{d \times d} \to \mathbb{R}$ is a GP-prior on $L^2(G)$, i.e.

$$\mathbb{E}F(x) = 0, \qquad \operatorname{Cov}(F(x), F(x')) = k(x, x'), \qquad x, x' \in \mathbb{R}^d.$$
(2)

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Setting $K := (k(X_i, X_j))_{i,j=1}^n \in \mathbb{R}^{n \times n}$ and $k(X, x) := (k(X_i, x))_{i=1}^n \in \mathbb{R}^n$, the posterior $\Pi(\cdot|X, Y)$ is given by the GP with mean and covariance function

$$x \mapsto k(X, x)^{\top} (K + \sigma^2 I_n)^{-1} Y$$

$$(x, x') \mapsto k(x, x') - k(X, x)^{\top} (K + \sigma^2 I_n)^{-1} k(X, x').$$

$$(3)$$

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Motivating problem.

The computation of $(K + \sigma^2 I)^{-1}$ has a computational complexity of $O(n^3)$, which becomes infeasable for large *n*.

Idea: Focussing on the posterior mean $k(X, x)^{\top}(K + \sigma^2 I_n)^{-1}Y$, iteratively solve $(K + \sigma^2 I_n)W = Y$ for the representer weights W.

• Consider a Bayesian updating scheme updating scheme with initial believes $W = (K + \sigma^2 I_n)^{-1} Y \sim N(0, (K + \sigma^2 I_n)^{-1}) =: N(w_0, \Gamma_0).$

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Consecutively update by conditioning on the information projection

$$\alpha_j := \mathbf{s}_j^\top (\mathbf{Y} - (\mathbf{K} + \sigma^2 \mathbf{I}_n) \mathbf{w}_{j-1}), \qquad j \le m \tag{4}$$

where $s_i, j \leq m$ are search directions chosen by the user.

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Consider a Bayesian updating scheme updating scheme with initial believes W = (K + σ²I_n)⁻¹Y ~ N(0, (K + σ²I_n)⁻¹) =: N(w₀, Γ₀).

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• After *m* steps, believes are given by $N(w_m, \Gamma_m) = N(C_m Y, (K + \sigma^2)^{-1} - C_m)$. This yields the approximate Gaussian posterior $\Psi_m := \mathbb{P}^{F|W=w}N(w_m, \Gamma_m)(dw)$ with mean and covariance functions

$$x \mapsto k(X, x)^{\top} C_m Y \qquad (x, x') \mapsto k(x, x') - k(X, x)^{\top} C_m k(X, x'), \qquad (5)$$

where C_m is a rank *m* matrix approximating $(K + \sigma^2 I_n)^{-1}$.

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The Empirical eigenvector posterior

Consider the spectral decomposition of the empirical kernel matrix

$$K = \sum_{j=1}^{n} \widehat{\mu}_j \widehat{u}_j \widehat{u}_j^{\top}$$
(6)

and choose the search directions $s_j := \widehat{u}_j, \, j \leq m$.

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and choose the search directions $s_j := \widehat{u}_j$, $j \leq m$. Then, the approximate posterior $\Psi_m = \Psi_m^{\rm EV}$ is given by the mean and covariance function

$$x \mapsto k(X,x)^{\top} C_m Y \qquad (x,x') \mapsto k(x,x') - k(X,x)^{\top} C_m k(X,x'), \tag{7}$$

where $(K + \sigma^2 I_n)^{-1}$ is approximated by

$$C_m = C_m^{\mathsf{EV}} = \sum_{j=1}^m (\widehat{\mu}_j + \sigma^2)^{-1} \widehat{u}_j \widehat{u}_j^\top.$$
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The Ψ_m^{EV} is equivalent to the Variational Bayes posterior based on spectral inducing variables $[\text{NSZ22}]^2$

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$$K = \sum_{j=1}^{n} \widehat{\mu}_j \widehat{u}_j \widehat{u}_j^{\top}$$
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and choose the search directions $s_j := \tilde{u}_j, j \leq m$, where $(\tilde{\mu}_j, \tilde{u}_j)_{j \leq m}$ is the Lanczos approximate eigensystem up to order m.

$$\mathcal{K} = \sum_{j=1}^{n} \widehat{\mu}_{j} \widehat{u}_{j} \widehat{u}_{j}^{\top}$$
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$$x \mapsto k(X,x)^{\top} C_m Y \qquad (x,x') \mapsto k(x,x') - k(X,x)^{\top} C_m k(X,x'), \tag{10}$$

with

$$C_m = C_m^{\rm L} = \sum_{j=1}^m (\tilde{\mu}_j + \sigma^2)^{-1} \tilde{u}_j \tilde{u}_j^{\rm T}.$$
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Randomness of the kernel matrix

Since $K = (k(X_i, X_j)_{i,j \le n})$ is a random matrix, the spectral decomposition of K cannot be computed in advance.

$$\varrho(w_j) = \min_{t \in \mathbb{R}} \varrho(w_{j-1} + td_j^{\mathsf{CG}}), \tag{12}$$

where $\varrho(w) := (w^{\top}(K + \sigma^2 I_n)w)/2 - Y^{\top}w$, and the $(d_j^{CG})_{j\geq 1}$ are conjugate search directions satisfying $(d^{CG})_i^{\top}(K + \sigma^2 I_n)d_k^{CG} = 0, j \neq k$.

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For the policies $s_j := d_j^{CG}$, $j \le m$, Bayesian updating is equivalent to the CG-iteration and we obtain the approximate posterior Ψ_m^{CG} given by

$$x \mapsto k(X, x)^{\top} C_m Y \qquad (x, x') \mapsto k(x, x') - k(X, x)^{\top} C_m k(X, x'), \tag{13}$$

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GPU accelerated matrix vector multiplication

CG only relies on matrix vector multiplications, which can be GPU accelerated and makes CG particularly relevant for large scale applications, see Wang et al. [Wan+19].

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Reduction in computational complexity

The approximate inversions C_m^L , C_m^{CG} have a computation cost of $O(mn^2)$, which is feasible when $m \ll n$.

Theorem (Approximate posterior contraction)

For $f_0 \in \overline{\mathbb{H}}$ with $\mathbb{H} = \operatorname{ran} T_k^{1/2}$,

$$T_k: L^2(G) \to L^2(G), \qquad f \mapsto \int f(y)k(\cdot, y) G(dy) = \sum_{j=1}^{\infty} \lambda_j \langle f, \phi_j \rangle_{L^2(G)} \phi_j, \qquad (14)$$

let \mathbb{P}_{f_0} be the measure corresponding to the data generating process

$$Y_i = f_0(X_i) + \varepsilon_i, \qquad i = 1, \dots n \tag{15}$$

Then, the true posterior Π_n satisfies that for any sequence $M_n \to \infty$,

$$\Pi_n(\{f \in L^2(G) : d(f, f_0) \ge M_n \varepsilon_n\} | X, Y) \to 0$$
(16)

in probability under $\mathbb{P}_{f_0}^{\otimes n}$ and $n \to \infty$, where ε_n is the optimal achievable rate implied by a concentration function inequality.

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Then, the approximate posterior Ψ_m satisfies that for any sequence $M_n \to \infty$,

$$\Psi_{m_n}(\{f \in L^2(G) : d(f, f_0) \ge M_n \varepsilon_n\} | X, Y) \to 0$$
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in probability under $\mathbb{P}_{f_0}^{\otimes n}$ and $n \to \infty$, where ε_n is the optimal achievable rate implied by a concentration function inequality and $m_n \to \infty$ is an appropriate sequence. For an ONB $(\phi_j)_{j\geq 1}$ of $L^2(G)$ and $Z_j \sim N(0,1)$ i.i.d., consider the random series prior

$$F(x) = \sum_{j=1}^{\infty} \tau j^{-1/2 - \alpha/d} Z_j \phi_j(x), \qquad x \in \mathbb{R}^d$$
(17)

where $\alpha >$ 0 and τ are the regularity and scale hyperparameters of the process. Then, for any

$$f_0 \in S^{\beta}(L) := \{ f \in L^2(G) : \|f\|_{S^{\beta}}^2 \le L \} \quad \text{with} \quad \|f\|_{S^{\beta}}^2 := \sum_{j=1}^{\infty} j^{2\beta/d} \langle f, \phi_j \rangle^2, \quad (18)$$

with $d/2 < \beta \le \alpha + d/2$ and an apropriate choice of τ , the approximate posterior satisfies that for any $M_n \to \infty$,

$$\Psi_{m_n}\{f: d_{\mathsf{H}}(f, f_0) \ge M_n n^{-\beta/(d+2\beta)} | X, Y\} \to 0,$$
(19)

in probability under $\mathbb{P}_{f_0}^{\otimes n}$ and $n \to \infty$ with $m_n \sim n^{d/(2\beta+d)} \log n$.



Figure 1: Simulation results for n = 3000, m = 20, 40.



Figure 2: Simulation results for n = 3000, m = 80 and scaling of computation times.

- Our theory is the first providing statistical guarantees for fully numerical algorithms.
- Particular relevance in the CG posterior. Default method in the GPyTorch library, see Gardner et al. [Gar+18].

- ► For KL(Ψ_{m_n} , $\Pi_n(\cdot|X, Y)$) $\leq n\varepsilon_n^2$, the approximate posterior Ψ_{m_n} inherits the contraction rate ε_n , see Ray and Szabó [RS19].
- For the empricial eigenvector posterior with $s_j = \hat{u}_j$, $j \le m$ this bound is available via elementary tools.
- Analyze the Lanczos posterior as an approximation.

Theorem (Lanczos: Eigenvalue bound, [Saa80])

Under Assumption (LWdf), for any fixed integer $i \leq \tilde{m} < n$ with $\tilde{\lambda}_{i-1} > \hat{\lambda}_i$ if i > 1and any integer $\tilde{p} \leq \tilde{m} - i$, the eigenvalue approximation satisfies

$$0 \leq \widehat{\lambda}_{i} - \widetilde{\lambda}_{i} \leq (\widehat{\lambda}_{i} - \widehat{\lambda}_{n}) \Big(\frac{\widetilde{\kappa}_{i} \kappa_{i, \tilde{p}} \tan(\widehat{u}_{i}, v_{0})}{T_{\tilde{m} - i - \tilde{p}}(\gamma_{i})} \Big)^{2},$$
(20)

where $\gamma_i := 1 + 2(\widehat{\lambda}_i - \widehat{\lambda}_{i+\widetilde{p}+1})/(\widehat{\lambda}_{i+\widetilde{p}+1} - \widehat{\lambda}_n)$,

$$\tilde{\kappa}_i := \prod_{j=1}^{i-1} \frac{\tilde{\lambda}_j - \hat{\lambda}_n}{\tilde{\lambda}_j - \hat{\lambda}_i}, \qquad \kappa_{i,\tilde{\rho}} := \prod_{j=i+1}^{i+\tilde{\rho}} \frac{\hat{\lambda}_j - \hat{\lambda}_n}{\hat{\lambda}_i - \hat{\lambda}_j},$$
(21)

and T_I denotes the I-th Tschebychev polynomial.

Challenges from spectral concentration

Theorem (Eigenvalue concentration, Shawe-Taylor and Williams [STW02]) The empirical eigenvalues $(\hat{\lambda}_j)_{j \le n}$ of the normalized kernel matrix K/n satisfy (i) For any t > 0 and any fixed m > 1, both

$$\mathbb{P}\{|\widehat{\lambda}_m - \mathbb{E}\widehat{\lambda}_m| \ge t\} \le 2\exp\left(\frac{-2nt^2}{\max_x k(x,x)^4}\right)$$
(22)

and

$$\mathbb{P}\{|\sum_{j=m+1}^{n}\widehat{\lambda}_{j} - \mathbb{E}\sum_{j=m+1}^{n}\widehat{\lambda}_{j}| \ge t\} \le 2\exp\Big(\frac{-2nt^{2}}{\max_{x}k(x,x)^{4}}\Big).$$
(23)

(ii) For any fixed $m \ge 1$,

$$\mathbb{E}\sum_{j=1}^{m}\widehat{\lambda}_{j} \ge \sum_{j=1}^{m}\lambda_{j} \quad \text{and} \quad \mathbb{E}\sum_{j=m+1}^{n}\widehat{\lambda}_{j} \le \sum_{j=m+1}^{\infty}\lambda_{j}. \quad (24)$$

Proposition (Relative perturbaton bounds, Jirak and Wahl [JW23])

Under appropriate assumptions, fix $m \in \mathbb{N}$ and further assume that

$$\mathbf{r}_{j}(T_{k}) := \sum_{k \neq j} \frac{\lambda_{k}}{|\lambda_{j} - \lambda_{k}|} + \frac{\lambda_{j}}{(\lambda_{j-1} - \lambda_{j}) \wedge (\lambda_{j} - \lambda_{j+1})} \leq C \sqrt{\frac{n}{\log n}}, \qquad (22)$$

for all $j \leq m$. Then, the eigenvalues of K/n satisfy the relative perturbation bound

$$\frac{\widehat{\lambda}_{j} - \lambda_{j}}{\lambda_{j}} \Big| \le C \sqrt{\frac{\log n}{n}} \qquad \text{for all } j \le m$$
(23)

with high probability.

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Martin Wahl

Ongoing joint work on perturbation series for empirical eigenvalues and eigenprojectors.

- ► For KL(Ψ_{m_n} , $\Pi_n(\cdot|X, Y)$) $\leq n\varepsilon_n^2$, the approximate posterior Ψ_{m_n} inherits the contraction rate ε_n , see Ray and Szabó [RS19].
- ▶ For the empricial eigenvector posterior with $s_j = \hat{u}_j$, $j \leq m$ this bound is available via elementary tools.
- Analyze the Lanczos posterior as an approximation.
- Establish the equivalence of the CG and the Lanczos posterior.

- Our theory is the first providing statistical guarantees for fully numerical algorithms.
- Particular relevance lies in the CG posterior. Default method in the GPyTorch library, see Gardner et al. [Gar+18].
- New interpretation of the CG posterior as a numerical approximation of a variational Bayes method.



Thank you!

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